On the use of predicted values in nutritional epidemiology: Example from the Hispanic Community Health Study/Study of Latinos (HCHS/SOL)

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Introduction

- In epidemiology, there are many measurements that are difficult to obtain directly:
 - Expensive (Resting Energy Expenditure)
 - Burdensome (24 hour urinary sodium)
 - Impossible (Usual energy intake)
- One strategy is to use prediction equations to measure them indirectly
- Many analyses proceed with predicted values as if they were observed data
- Using predicted values instead of observed data in study analyses can corrupt study results if the (Berkson) prediction error is not handled appropriately

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Berkson vs Classical measurement error (Keogh et al 2021)

Classical error adds random noise to the true value X X* = X + error

Example: A single measure of blood pressure X* can fluctuate randomly around an innate true average value X

Observations with Berkson error are less variable than true value X
 X = X* + error

Example: A predicted value from a regression equation has less variability than the original outcome, due to unexplained variance

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Example from the Hispanic Community Health Study (Lavange et al 2010)

Question of interest: Does sodium intake vary by Hispanic ethnicity?

HCHS main cohort: n = 16,415

Male: 40%

Age: mean 43y; range: 18-74y

Main dietary assessment: two 24 hour recalls, known to be subject to bias

SOLNAS: Calibration sub-study: n = 477

Biomarker: 24 hour urinary sodium was obtained to create calibration equations that correct for the measurement error/bias in self-reported sodium (Mossavar-Rahmani et al 2017)

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Calibration equations as prediction equations

If a biomarker Y** has classical error one can estimate true intake (Y)

by regressing Y^{**} on self-reported Y^{*} and other covariates (age, BMI, gender, language preference, restaurant score, fast food intake)

<u>Step 1</u>: use Y** to Fit Model:

 $Y = b_0 + b_1 Y^* + b_2 X^2 + b_3 X^3 + \dots + b_k X^k + epsilon$

Step 2: Use fitted regression equation to derive predicted (mean) intake for a give sent of covariates.

•
$$\hat{Y} = \hat{b}_0 + \hat{b}_1 Y^* + \hat{b}_2 X2 + \hat{b}_3 X3 + \dots \hat{b}_k Xk$$

• The unexplained variance from the calibration equation results in the Berkson error in measure \hat{Y}

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$$Y = \hat{Y} + e$$

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A simple fix for Berkson error

- The fundamental problem of predicted values is their Berkson error makes them less variable than they should be.
- When the distribution is of interest, a simple fix is to add back the missing variance to the calibrated value.
 - This can be accomplished from simulating error e ~ N(0, σ^2), where $\sigma^2 = \sigma^2_{Resid}$ σ^2_{within}
 - $Y_{imp} = \hat{Y} + \mathbf{e}$

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- A multiple imputation approach can be used to estimate the quantities of interest
- If the error-prone variable Y* only had classical measurement error, the NCI method (SAS macro) could be used to estimate the distribution (Tooze et al 2006)

Numerical Study

Here consider a computer simulation study

- $X_i = \gamma_0 + \gamma_1 X_i^* + \gamma_2 Z_i + \varepsilon_i$ (relationship between truth and self-report)
- $X_i^{**} = X_i + u_i$. (relationship between truth and biomarker)

Fit calibration equation and compare distributions of

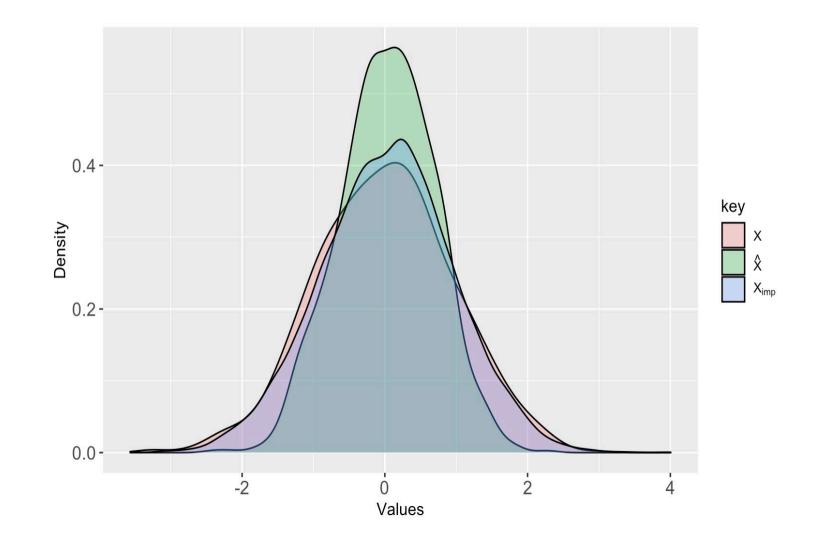
<u>True X</u> <u>Predicted values</u> \hat{X} : From calibration equation regressing X_i^{**} on (X,Z) <u>Imputed X</u>_{imp}: $X_{imp,i}^{(m)} = \hat{X}_i^{(m)} + e_i^m$

Simulation settings: normal distributions (simplistic); classical error in X*; var(ε)=var(X); 1000 simulations; 1000 bootstraps (CI for Ximp quantiles) R software (version 4.04)

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Simulation Study Results



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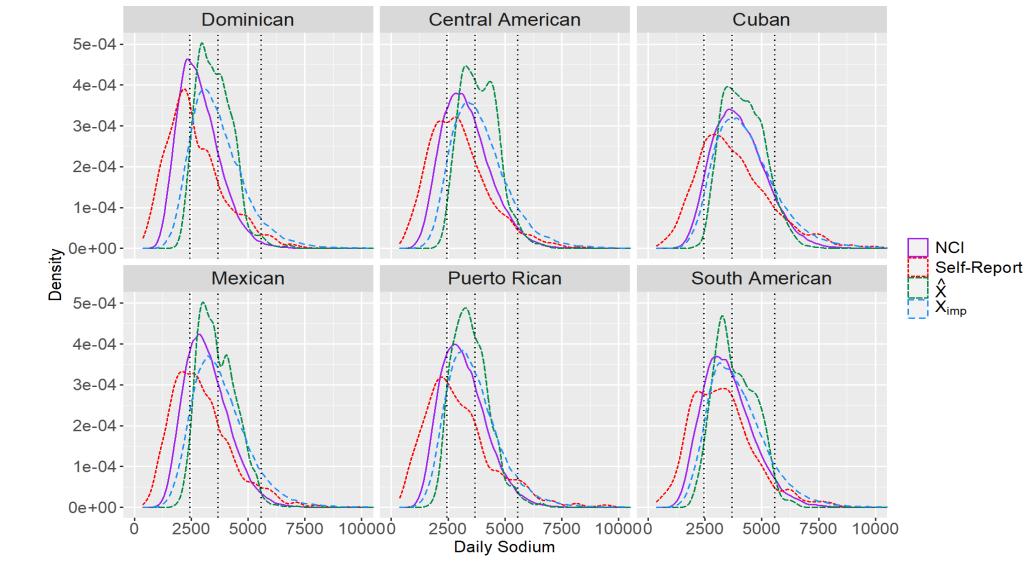
Berkson error biases quantiles and standard errors

| | | Χ | | | Â | | | X _{imp} | |
|--------|--------|-------|------|--------|-------|-----|--------|------------------|------|
| %-tile | Mean | ESE | CP | Mean | ESE | CP | Mean | ESE | CP |
| 10th | -1.277 | 0.055 | 94.7 | -0.949 | 0.105 | 1.3 | -1.289 | 0.121 | 97.1 |
| 25th | -0.672 | 0.043 | 94.8 | -0.501 | 0.079 | 8.0 | -0.679 | 0.089 | 96.6 |
| 50th | -0.001 | 0.039 | 96.1 | -0.002 | 0.067 | 6.0 | -0.002 | 0.074 | 97.4 |
| 75th | 0.674 | 0.043 | 94.5 | 0.498 | 0.078 | 8.3 | 0.675 | 0.087 | 96.5 |
| 90th | 1.276 | 0.053 | 95.8 | 0.947 | 0.103 | 0.8 | 1.284 | 0.119 | 96.9 |



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Similar results seen in HCHS/SOL



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Marked differences in percentiles and SE (low R-squared)

| Percentile | Self-report X* | Treated as observed | X _{imp} |
|------------|----------------|---------------------|------------------|
| 10th | 1571 (66) | 2711 (37) | 2401 (164) |
| 25th | 2091 (63) | 3017(33) | 2869 (125) |
| 50th | 2853 (80) | 3514 (56) | 3547 (102) |
| 75th | 3920 (108) | 4153 (77) | 4422 (168) |
| 90th | 5026 (132) | 4837 (78) | 5371 (323) |

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Proportion with Sodium Intake <2300 mg/day

| Ethnic Background | Self-Report (X*) | X (treated as observed) | X _{imp} (account for Berkson error) |
|---------------------------|------------------|-------------------------|---|
| Dominican (n= 1451) | 0.456 (0.018) | 0.025 (0.005) | 0.104 (0.055) |
| Central American (n=1708) | 0.334 (0.018) | 0.004 (0.002) | 0.054 (0.034) |
| Cuban (n=2329) | 0.174 (0.009) | 0.001 (0.001) | 0.034 (0.023) |
| Mexican (n=6405) | 0.347 (0.010) | 0.018 (0.003) | 0.080 (0.044) |
| Puerto Rican (n=2677) | 0.367 (0.014) | 0.033 (0.004) | 0.111 (0.045) |
| South American (n=1061) | 0.272 (0.018) | 0.020 (0.007) | 0.067 (0.036) |
| All (N=15825) | 0.320 (0.007) | 0.016 (0.001) | 0.075 (0.004) |





Comments regarding other analyses

- Predicted values as covariates in a regression (classic regression calibration)
 - Berkson error in a covariate will not bias regression coefficient (so long as prediction equation correct)
 - Standard errors still need to be adjusted to account for uncertainty in predict model coefficients
- Predicted values as the outcome in a regression (classic regression calibration)
 - Need Berkson error to be independent of the covariates in the regression model
 - Coefficients will be biased
 - Buonaccorsi method (1991) can be used to address bias, so long as non-differential error in predicted value

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Discussion

- There is increasing use of prediction and calibration equations in medicine
- Naïve analyses with predicted outcomes are subject to multiple biases
 - Distributional summaries are biased, quantiles appear less extreme
 - Regressions reliant on predicted outcomes will have biased coefficients
 - Regressions reliant on predicted exposures need SE adjustment

Presented methods do not address when prediction error is differential

- Deficiencies in the prediction model leads to correlation between prediction error and other analysis variables
- Recent work (Haber et al ; Ogburn et al 2021) has outlined bias related to misspecified prediction models
- Awareness of the effects of Berkson error and methods to adjust for it need more attention

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