Cautionary notes for regression analyses that use predicted values as an outcome or exposure

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Introduction

- In epidemiology, there are many measurements that are difficult to obtain directly:
 - Expensive (Resting Energy Expenditure)
 - Burdensome (24 hour urinary sodium)
 - Impossible (Usual energy intake)
- One strategy is to use prediction equations to measure them indirectly
- Many analyses proceed with predicted values as if they were observed data
- Using predicted values instead of observed data in study analyses can corrupt study results if the (Berkson) prediction error is not handled appropriately

Berkson vs Classical measurement error (Keogh et al 2021)

Classical error adds random noise to the true value X X* = X + error

Example: A single measure of blood pressure X* can fluctuate randomly around an innate true average value X

Observations with Berkson error are less variable than true value X
 X = X* + error

Example: A predicted value \hat{X} from a regression equation has less variability than the original outcome, due to unexplained variance

Example from the Hispanic Community Health Study/Study of Latinos

(Lavange et al 2010)

Questions of interest: Is potassium intake associated with hypertension? Does potassium intake vary by level of acculturation or Hispanic ethnicity?

HCHS/SOL main cohort: N = 16,415, recruited 2008-2011 from the Bronx, Chicago, Miami and San Diego

Male: 40%

Baseline Age: mean 43y; range: 18-74y

Dietary assessment: two 24 hour recalls, known to be subject to bias

SOLNAS: Calibration sub-study: n = 477

Biomarker: 24 hour urinary potassium was obtained to create calibration equations that correct for the measurement error/bias in self-reported sodium. A subset had repeated measures of biomarker (Mossavar-Rahmani et al 2017)

Regression Calibration

• Popular method for addressing covariate measurement error

Suppose:

$$Y = \beta_0 + \beta_X X + \beta_Z Z + \varepsilon$$
 and
 $X^* = \alpha_0 + \alpha_X X + \alpha_Z Z + U$

Then

$$E[Y|X^*, Z] = E_{X|X^*, Z}[E(Y|X^*, Z)|X] = E_{X|X^*, Z}[E(Y|Z, X)]$$
$$= E_{X|X^*, Z}[\beta_0 + \beta_X X + \beta_Z Z]$$
$$= \beta_0 + \beta_X E[X|X^*, Z] + \beta_Z Z$$

Conclusion: regress Y on $E[X|X^*, Z]$ and Z to get right β coefficients. $E[X|X^*, Z]$ is referred to as the calibrated exposure

Calibration equations as prediction equations

If a biomarker X^{**} has classical error one can estimate true intake (X)

by regressing X^{**} on self-reported X^{*} and other covariates (age, BMI, gender, language preference, restaurant score, fast food intake)

<u>Step 1</u>: use X^{**} to Fit Model:

 $X = b_0 + b_1 X^* + b_2 Z1 + b_3 Z2 + \dots b_{k+1} Zk + epsilon$

Step 2: Use fitted regression equation to derive predicted (mean) intake for a give sent of covariates.

•
$$\hat{X} = \hat{b}_0 + \hat{b}_1 X^* + \hat{b}_2 Z1 + \hat{b}_3 Z2 + \dots \hat{b}_{k+1} Zk$$

• The unexplained variance from the calibration equation results in the Berkson error in measure \hat{X}

•
$$X = \hat{X} + e$$

Predicted values as covariates in a regression

- Berkson error in a covariate will not bias a linear regression coefficient (so long as prediction equation correct, independent error)
- Approximation if non linear outcome model

Many common and underappreciated pitfalls when applying regression calibration

- Standard errors still need to be adjusted to account for extra uncertainty
- Prediction model needs all covariates in outcome model to avoid bias
- Extra covariates can be included in prediction model if not correlated with outcome given the truth
- Special considerations when calibration model covariate is a mediator

Simulation Study: Regression Calibration in logistic regression

Variables	Parameter values
$X^* = a_0 + a_1 X + a_2 Z + a_3 V + e$	$a_0 = 0.4, a_1 = 0.5, a_2 = 0.5, a_3 = 0.2;$
	$\sigma_{e}^{2} = 0.49$
$X^{**} = X + d$	$\sigma_{d}^{2} = 0.7$
(X,Z,V)	Multivariate normal
	$\underline{\mu_X} = \underline{\mu_Z} = \underline{\mu_V} = 0$
	$\underline{\sigma_X^2} = \underline{\sigma_Z^2} = \underline{\sigma_V^2} = 1$
	cor(X,Z) = cor(X,V) = cor(Z,V) = 0.5
$Logit(Y) = b_0 + b_1X + b_2Z + b_3V$	$b_0 = -1.0, b_1 = log(1.5), b_2 = -log(1.3), b_3$
	$=\log(1.75)$

Numerical Study

- Results from 1000 simulations of regression calibration:
- Cohort N=2500; calibration substudy n=250

Method	Mean	% Bias	Empirical standard error	Average estimated standard error	Coverage probability
Model-based Bootstrap- based	0.407	0.3	0.136	0.113 0.140	0.915

Underappreciated bias when models not aligned

Method	Mean	Empirical Standard Error of Mean	% Bias
Naïve regression	0.201	0.057	-50.3
Correct RC model	0.407	0.136	0.3
RC, Non-aligned outcome model	0.912	0.194	125.0
RC, Non-aligned calibration model	0.366	0.115	-9.7

Returning to HCHS Example

- Is potassium associated with lower odds of hypertension?
 - For the outcome model: also adjust for potential confounders: age, sex, Hispanic/Latino background, education, income, current smoking, body mass index (BMI)
 - Supplement intake is a useful covariate for the calibration model
 - Recommended approach: include supplement intake into both the outcome and calibration models.

HCHS Analysis: Results

Method of Estimation	OR	95% CI*
Including supplement use in outcome model	0.76	0.60 – 0.96
Omitting supplement use from outcome model	0.90	0.75 – 1.07

Regression with a predicted outcome

Measurement Error Model: $Y = \hat{Y} + e$

Outcome model of interest: $Y = \beta_x X + \beta_z Z + u$

Common Setting: Want to relationship between Y and X, but Y hard to measure. Prediction equation developed in previous study.

Fundamental Question: If fit model $\hat{Y} = \beta_x^* X + \beta_z^* Z + v$ will $\beta_x^* = \beta_x^*$ **Answer:** No Intuition: $\hat{\beta} = (X^T X)^{-1} \sigma^2$

Impact of Berkson error in the Y



Intake

Addressing bias from Berkson error in outcome

 $\beta_x^* = \operatorname{var}(Y_{\text{pred}})/\operatorname{var}(Y) \beta_x$, if non-differential measurement error in \hat{Y} , that is if $f(\hat{Y} \mid Y, X, Z) = f(\hat{Y} \mid Y, Z)$

• If non-differential error can apply Buonaccorsi (1991) adjustment:

 $Y_{adj} = (\hat{Y} - \alpha_0) / \alpha_1$ $\alpha_1 = \operatorname{var}(\hat{Y}) / \operatorname{var}(Y) \text{ and } \alpha_0 = \mu_{\hat{Y}} - \alpha_1 \mu_Y$

- Differential error can occur if X or other confounders should have been in prediction model for $\hat{Y}_{.}$
 - For linear regression example can test: $\hat{Y} \coprod X \mid Y, Z$
 - The challenge: May not have (X,Y,Z) all in same dataset
 - In data example, the objective biomarker M=Y+error can be used to estimate Buonaccorsi coefficients and test this condition
 - For Buonaccorsi adjustment, estimate: var(Y) = var(M) var(M1-M2)/2

Does sodium intake vary with acculturation (English preference)?

 Calibration (prediction) equation obtained from regression of In (M) on log 24h recall sodium (24hrK), age, BMI, gender

ln(NA) = 7.268 + 0.136 ln(24hrK) + 0.001 age + 0.017 BMI - 0.274 l(Female)

- Can use M to test non-differentiality condition, equivalent to $\hat{Y} \coprod X \mid Y, Z$ in example
- Check regression of \hat{Y} on X,Y, Z.
- Since M = Y + error, can perform second regression calibration.
 Replace Y with E(Y|M,X,Z)= E(M2|M1,X,Z)

Three ways of regressing sodium intake on acculturation (English preference)

- **1**. Regress \hat{Y} on acculturation (biased)
- 2. Buonaccorsi's correction (unbiased if error is non-differential)
- 3. Regress M on acculturation (unbiased) E(M|X,Z) = E(Y+e|X,Z) = E(Y|X,Z)
- Methods 1 and 2 can be done in full HCHS or SOLNAS
- Method 3 can be done only in SOLNAS
- Method 3 is usually not available but here the SOLNAS substudy makes it possible.

HCHS/SOL Results Regression coefficient and SE

	\widehat{Y}	Buonaccorsi Adjustment	Unbiased
SOLNAS	0.064 (0.026)	0.166 (0.085)	-0.056 (0.056)
HCHS	0.029 (0.017)	0.075 (0.051)	

 \hat{Y} method appears biased as expected, and Buonaccorsi adjustment appears to increase the bias!

- Differential error check: regress \hat{Y} on X,Z, E(Y|X,Z)
- Regression Coefficient for X = 0.235; 95%CI = (0.095, 0.711)

Other examples of differential error

Intake	Regression coefficient for English preference	95% CI
Potassium	0.125	(0.016, 0.531)
Protein	0.101	(0.043, 0.174)
Total energy	0.038	(-0.007, 0.083)

Discussion

- There is increasing use of prediction/calibration equations in medicine
- Naïve analyses with predicted outcomes are subject to multiple biases
 - Regressions reliant on predicted outcomes will have biased coefficients
 - Regressions reliant on predicted values need SE adjustment
 - Distributional summaries are biased, quantiles appear less extreme
- Prediction model needs to be correct or all bets are off
 - This includes alignment of outcome and prediction model covariates
- Presented methods do not address when prediction error is differential
 - Deficiencies in the prediction model leads to correlation between prediction error and other analysis variables
 - Recent work (Haber et al ; Ogburn et al 2021) has outlined bias related to misspecified prediction models
- Awareness of the effects of Berkson error and methods to adjust for it need more attention

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